3. STATISTICAL ANALYSIS

3.1. Statistics of the Conductivity data set.

The results of the measurements, after the mean Conductivity value for each point is calculated, form a data set of 80 values. The statistics of this data set are calculated, by running the computer program VARIO1. The program's modified version VARIO3 is situated in appendix A. The arithmetic mean was found to be 65.79 m/day, and the sample variance $1\cdot10^4 \, (\text{m/day})^2$. The coefficient of skewness was found equal to $1\cdot915$, which means that the distribution is skewed to the right. The coefficient of kurtosis is found 6.73 which is much greater than 3, showing a leptokurtic distribution.

The very high value of the variance, shows the big variability of hydraulic conductivity in the studied area. The lowest value observed was $3\cdot 10^{-4}$ m/day and the highest 488 m/day. This variability should be expected since observation points expand over an area of about 1 square kilometer.

From the above-mentioned moments of the data set, it can be concluded that the conductivity in the studied area is not normally distributed, which is to be further examined, since this information only gives a general estimate of the parameter.

Considering as parameter under study the logarithm of the conductivity (under base 10), we find the same moments for

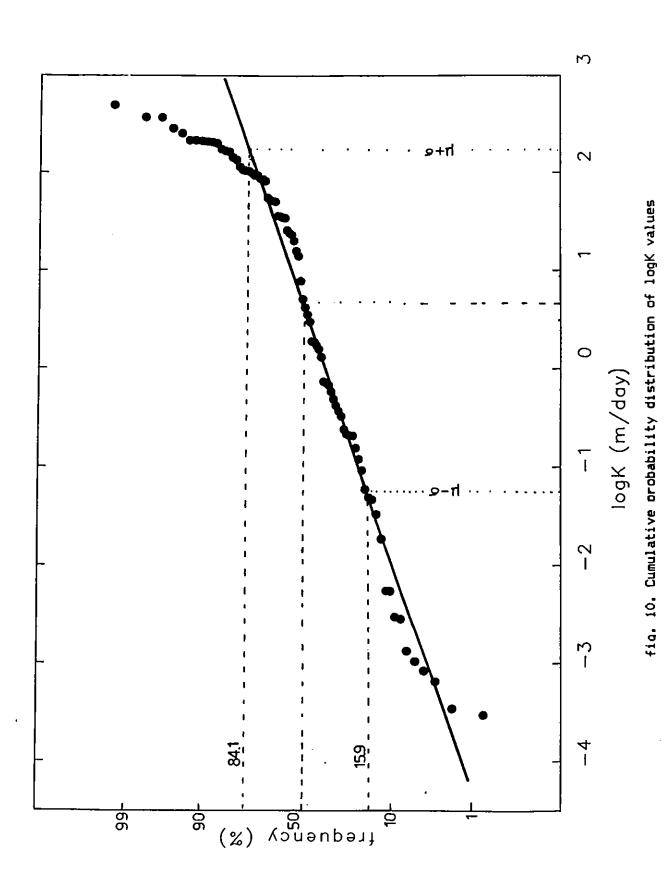
this parameter. The arithmetic mean is found equal to 0.491, with a sample variance of 3.062 (standard deviation 1.75). The skewness coefficient is -0.696 and the coefficient of kurtosis 2.34. Hence, the experimental distribution is found to be sligthly platukurtic and skewed to the left.

A lognormal distribution might fit the data. To get an idea about that, we plot the cummulative relative frequencies, expressed as percentages, against the observations (logK values), in a special probability graph paper (figure 10). The theoretical lognormal distribution, with the same mean and variance as the data, is also shown in the figure. The curve fits the data rather well. The above-mentioned graph, gives only an idea about the distribution, since no statistical test is possible to verify the fit, because the data are spatially correllated. The statistical test of Kolmogorov-Smirnov will be applied in a later chapter.

3.2. Comparison of results with previous studies.

As it was mentioned before, two different studies had taken place in the same region. The second one, timewise (Tan 1986), had a plot of 90x90 meters, situated near the South-East corner of the present study. It included observations of 100 points situated on the nodes of a grid with distances of 10 m.

The first one (Nurul 1984), had a plot of 14x14 meters, which was included in the above-mentioned plot. The samples were



0.000

taken at the nodes of a grid, every 2 meters.

27

In figure 11, the cumulative distributions of the three data sets are shown. One can observe there the gradual increase of the variance of the data set, with the increase of the area. The variance is proportional to the slope of the line of each data set, with respect to the normal distribution axis. The variance increase is better shown in figure 12, where the observations of each data set, have been brought around their mean.

The influence of the study area size on the variance was investigated by plotting the three variances versus the logarithms of the area (figure 13). In the figure it can be seen that a parabolic or exponential relationship appears to exist, unlike the linear relationship expected (A. G. Journel and Ch. J. Huijbregts, 1978).

The statistics of the three studies are compared and shown in table 8. The mean of the new data set, which corresponds to an area much larger than the other ones, is much bigger than the other means. The ratioes of the variances of the new data set with the "older" ones, are also very big.

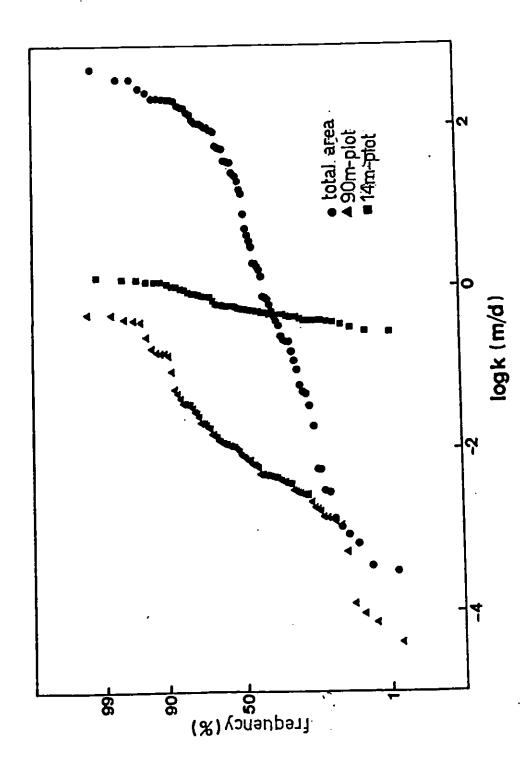
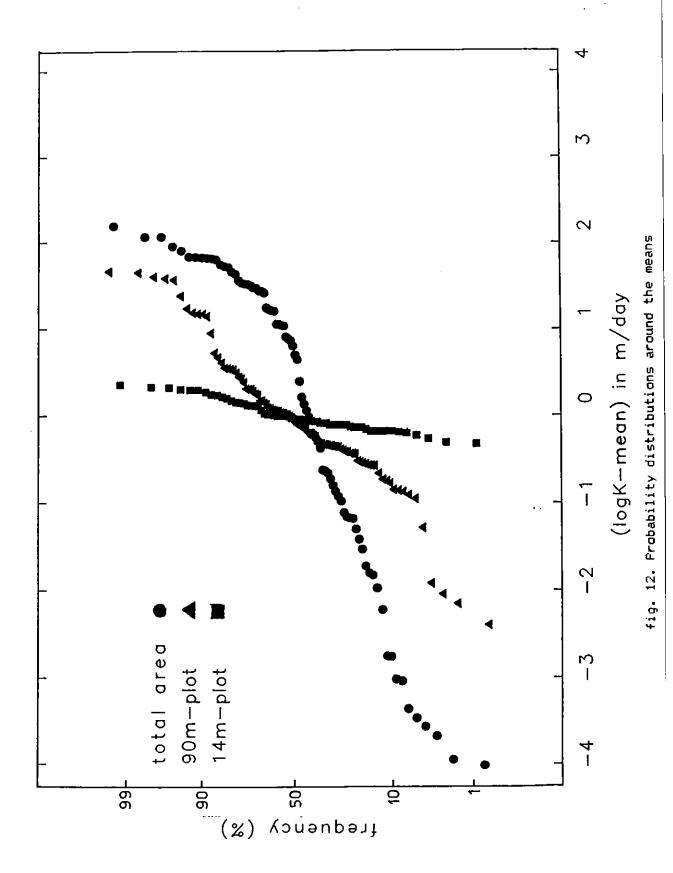
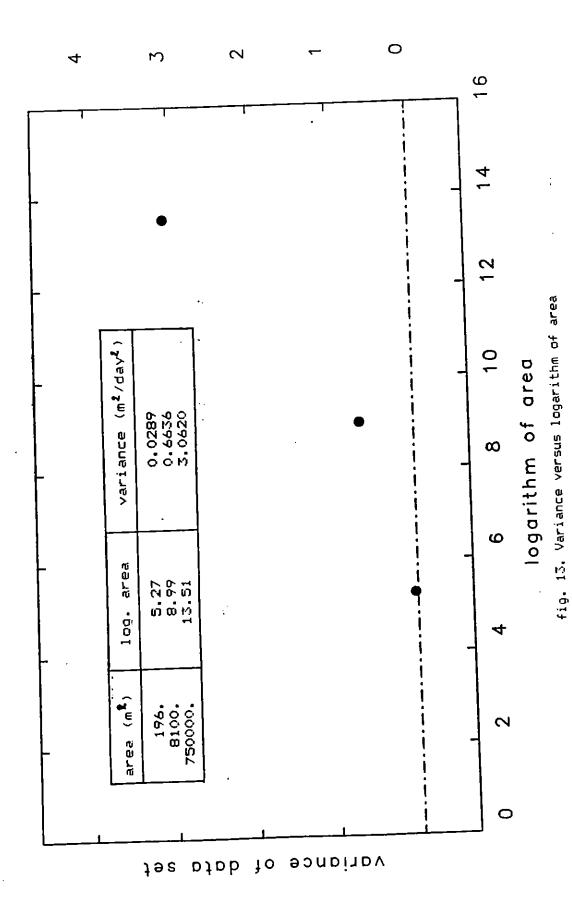


fig. 11. Cumulative distribution of the three studies





: 3

Table 8. Comparison of the statistics with previous studies.

Data set :		Nurul	ian	Present	fre./Nur.	Pre./Tan
s ean	1	0.6155	0.0404	65.788	106.9	1628.4
	2	-0.2451	-2.0070	0.491	-2.0	-0.2
variance	1	0.0688	Q. 00B9	10090.000	146721.0	1133707.9
	2	0.0289	0.6636	3.062	106.1	4.6
st. deviat.	1	0.2622	0.0943	100.455	383.1	1065.3
	2	0.1700	0.8146	1.750	10.3	2.1
SKENNESS	1	1.0480	3.2050	1.915	1.8	0.6
	2	0.4493	0.0589	-0.697	-1.6	-11.8
kurtosis	1	3. 0300	12.6000	6.728	2.2	0.5
	2	2.2370	3.9750	2.339	1.0	0.6

Note: 1 stands for K values, 2 for logK values

4. SPATIAL VARIABILITY

4.1. Introduction

Spatial variability analysis is the study of the differences that might exist between the value of a certain variable at a point with another at some distance in the same field area. Geostatistics was introduced in the middle of this century for mining purposes and geological studies. More recently, the technique has been applied in water resourses problems (Delfiner and Delhomme, 1973; Delhomme, 1978). Since then, many papers have been published about the study of spatial variability by the Geostatistical method applied to hydrogeology (Byers and Stephens, 1983; De Marsily, 1984; Gutjahr and Gelhar, 1981; Virdee and Kottegoda, 1984).

4.2. Geostatistics.

Geostatistics can be applied in the study of any phenomenon which can be characterised as a "regionalised phenomenon". It is called as such, a phenomenon that spreads out in space and shows certain structure. A variable which characterises such a phenomenon is termed as a regionalised variable (ReV). In fact all variables that describe properties of the subsurface or the atmosphere, may be considered as

regionalised variables.

From both conceptual and practical stand points, it is more convenient to deal with regionalised variables by applying the probabilistic theory of random functions (RP). But it is necessary to reconstitute the distribution law of this random function from the available data. The problem that arises here is that many regionalised variables have a unique existence, that is, one realisation. It is possible to compute the structure based only on this single outcome.

4.3. Hypotheses of stationarity and intrinsic.

Due to the above-mentioned problem, it is necessary to impose further hypotheses about the random function, so that it could be overcome.

The first hypothesis which is usually used in the theory of random functions, is the hypothesis of stationarity. This means that the expectation of the random function Z(x), is constant in space:

$$E(Z(x))=m \qquad (4.1)$$

x being the location vector.

The covariance depends only on the separation vector h: $cov(Z(x+h)+Z(x))=E(Z(x+h)-m)(Z(x)-m)=C(h) \qquad (4.2)$

The concept of ergodicity implies that the unique realisation will behave in space with the same probability

density function. In other words, by observing the variation in space of the property, it is possible to determine the probability density function of the random function for all realisations, which is termed as statistical inference.

In many natural phenomena, a finite variance does not exist. This leads to the introduction of the intrinsic hypothesis: Any increment of (Z(x+h)-Z(x)), has a finite variance which is independent of x:

$$E(Z(x+h)-Z(x))=0$$
 (4.3)
 $Var(Z(x+h)-Z(x))=2\chi(h)$ (4.4)

Equation (4.4) defines the variogram, which is a structural function of the parameter under study, and will be examined in the following in a higher extend.

Another hypothesis, the one of quasi-stationarity, is also necessary. It comes out of the fact that the mean, covariance and variogram are, in practise, functions of the size of the investigated area, as well as of the area itself. The structural function, covariance or variogram, is only used for limited distances [hkb. The limit b can represent for example, the diameter of the neighbourhood of estimation, or, in other cases, the extend of an homogeneous zone. It is allowed for every study area separately to use the locally found mean, covariance and variogram. However, they have no effect outside the area for which they have been calculated.

4.4. Variograms.

The variogram (h) is defined from the equation:

$$2\chi(h)=E(Z(x+h)-Z(x))$$
 (4.5)

It gives the mean squared difference in value for all pairs of measurements separated by a distance h. The term semivariogram is usually abreviated as "variogram", which may cause confusion to the readers. In the present paper, the term "variogram" will be used, unless otherwise stated.

The variogram considered, refers to point variables. Such a point variogram can be estimated from point measurements Z(x), as follows:

$$\chi(h) = (1/(2*N(h)) \sum_{i=1}^{N(h)} (Z(x+h)-Z(x))^{2}$$
 (4.6)

where N(h) is the number of pairs of samples separated by the vector h.

It can be proved that, if the regionalised variable is stationary, the relation between variance and covariance is:

$$\chi(h)=C(0)-C(h)$$
 (4.7)

where C(0) is the covariance at the point itself, and

C(h) is the covariance between points with distance h.

Variograms exhibit certain characteristics with the variation of distances. When the variogram has an increasing function, it starts from the origin, and increases until a certain value C_4 , which is the sill and is equal to the sample variance. The distance at which variograms reach the sill, is

called the range a. In physical sence, it expresses the fact that beyond this range, the samples are not correllated.

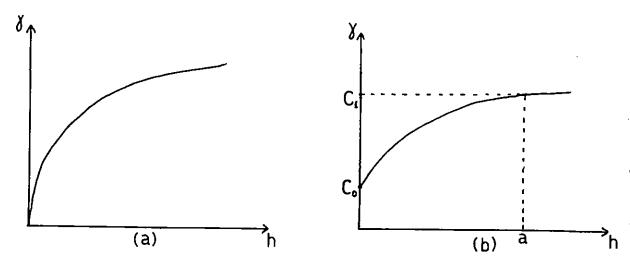


fig. 14. Variogram with sill, range, and nugget

In most cases, the variograms show a discontinuity close to the origin of the axes, which can be mathematically expressed by: $\lim_{n \to \infty} (\gamma(n)) = C_0 \qquad (4.8)$

 $C_{\rm o}$ is called the nugget effect. It represents the non-structural variability which can be spatial, or due to measurement errors.

The computation of a variogram can be carried out in one particular orientation, e.g. in a North-South grid, or in an East-West grid. When the variograms obtained from different orientations are equal, this means that we have isotropic properties. If that is not the case, we have anisotropism, that is, different properties in different directions.

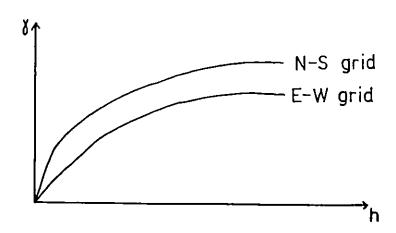


fig. 15. Variograms in different directions showing anisotr.

4.5. Variogram models.

It is usual practise to fix a mathematical description (model) to the variogram. The usual method for obtaining a mathematical description is to choose one of the proposed functional forms, and then calibrate it by using a statistical method, for instance, the least squares method.

The most commonly used models for variograms are:

Linear
$$\gamma(h) = C_0 + ah$$
 for $h < b$ (4.9a)
$$= C_0 + ab$$
 for $h > b$ (4.9b)
Spherical $\gamma(h) = C_0 + a(\frac{34}{2b} - \frac{4^3}{2b})$ for $h < b$ (4.10a)
$$= C_0 + a$$
 for $h > b$ (4.10b)
Power $\gamma(h) = C_0 + ah^b$ (4.11)
Logarithmic $\gamma(h) = C_0 + a\log(1+bh)$ (4.12)
Exponential $\gamma(h) = C_0 + a(1-\exp(-bh))$ (4.13)
Gaussian $\gamma(h) = C_0 + a(1-\exp(-bh^2))$ (4.14)

4.6. Computation of a variogram.

Each sample is considered as a point in the field.

Equation (4.6) is used. Hence, the contribution of the pair of points x1 and x2 which are in a distance h apart, is:

$$\gamma(h) = (1/2)(Z(x_i)-Z(x_i))^t$$
 (4.15)

As the computations for a set of data can be long and tedious, it is always easier to employ a computer for it. The program that was used here, is adopted from David, 1978 (refer to the appendices).

Algorithm of the program.

The main step is to sort all pairs of points available to a particular class, with respect to direction and distance. The classification pattern for a given direction, is as shown in figure 16.

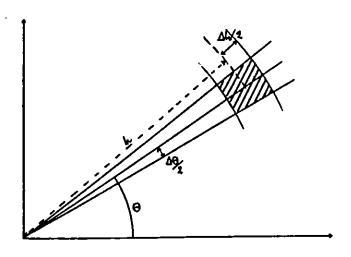


fig. 16. Classification pattern for the variogram computation

In the program, the direction θ is referred to as PHI and $\Delta\theta$ is referred to as PSI (the angular regularisation). Directional classification of the line joining two sample points x_4 and x_2 , is done by computing the scalar product

$$s = (\overline{x1x2})/|\overline{x1x2}| *\vec{u} \qquad (4.16)$$

where \vec{u} is the unit vector of the direction selected.

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The distance $x_1 \rightarrow x_2$ is classified according to $\triangle h$ and its combination.

$$C_{x_4x_2} = (Z(x_4) - Z(x_2))^2$$
 (4.17)

When all pairs have been tested and grouped against a certain direction class, the smoothed or average variogram along the direction, is expressed by:

$$\chi(h,\theta) = (1/2n_i) \sum_{x_i} (Cx_i x_2)$$
 (4.18)

where $h=\sum_{n}|x_{1}x_{2}|/n_{i}$, and $n_{i}=$ number of pairs grouped in the same class (i).

What is finally obtained is a set of one-dimensional smoothed variogram values along a selected direction. The calculation of the drift:

$$D(h_i, \theta) = (1/n_i) (\Sigma(x_k) - Z(x_\ell))$$
 (4.19)

enables us to detect the presence of trends.

The procedure of using the program VARIO1 is given in Appendix A.

4.7. Results.

The variograms of the K values data set and of the logK values data set were calculated. Originally, orientation PHI is taken 45 degrees and the angular regularisation PSI equal to 180 degrees, in order to have all possible pairs.

The maximum distance until which the variograms are calculated, is 500 m. Intervals of 100 meters were used. The results of the computations are situated in tables 9 and 10, and the respective plotts in figures 17 and 18.

The computations were done by using the PORTRAN program VARIO1 (M. David, 1977), with some alterations. The output includes some statistical parameters, and, for each class, the number of pairs, the drift and variogram values, the average pair distance, and the "maxvar pair", that is, the pair with the maximum contribution to the variogram value of the specific class. The results will be further discussed and explained, after examining the isotropy of the parameter.

4.7.1. Directional analysis.

In order to check whether the parameter (logK) has the same statistical properties in different directions, we calculate the variograms for the directions of 0, 45 and 90 degrees, with angular regularisation of 5 and 20 degrees. The results can be seen in tables 11 and 12, and their graphs in figures 19 and 20. In the first graph (5°), the variograms are very irregular for

table 9. The variogram of the K data set

VARIOORAM

STUDY OF SPATIAL VARIABILITY OF FIELD PARAMETER:
HYDRAULIC CONDUCTIVITY IN M/DAY
(WITH A FIELD OF 180. DEGREES IN EACH DIRECTION)

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							DRIFT	
	. 1000E+03	. 4883E+03	. 6579E+02	. 1009E+03	. 1915E+01	. 672BE+01	NO. OF PAIRS	120 273 461 461 276 276 276 271 271
			•	•		1	N METER	
	STEP IN METER	UPPER LIMIT FOR Z	GENERAL MEAN OF 2	GENERAL VARIANCE OF Z	GENERAL SKEWNESS OF Z	GENERAL KURTOSIS OF Z	DISTANCE IN METER	8.588888888888888888888888888888888888

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table 10. The variogram of the logK data set

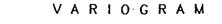
VARIOORAM

STUDY OF SPATIAL VARIABILITY OF FIELD PARAMETER: HYDRAULIC CONDUCTIVITY IN M/DAY (WITH A FIELD OF 180. DEGREEB IN EACH DIRECTION)

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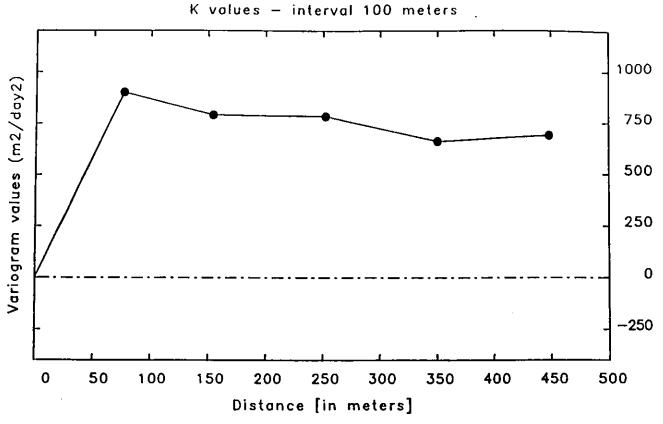


fig. 17. Variogram of the K data set

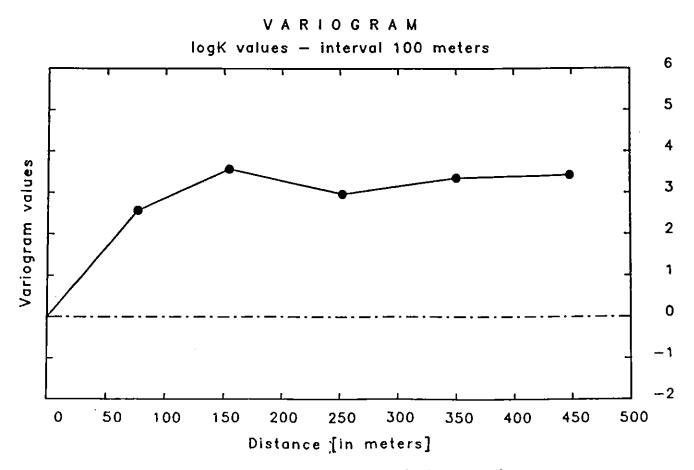


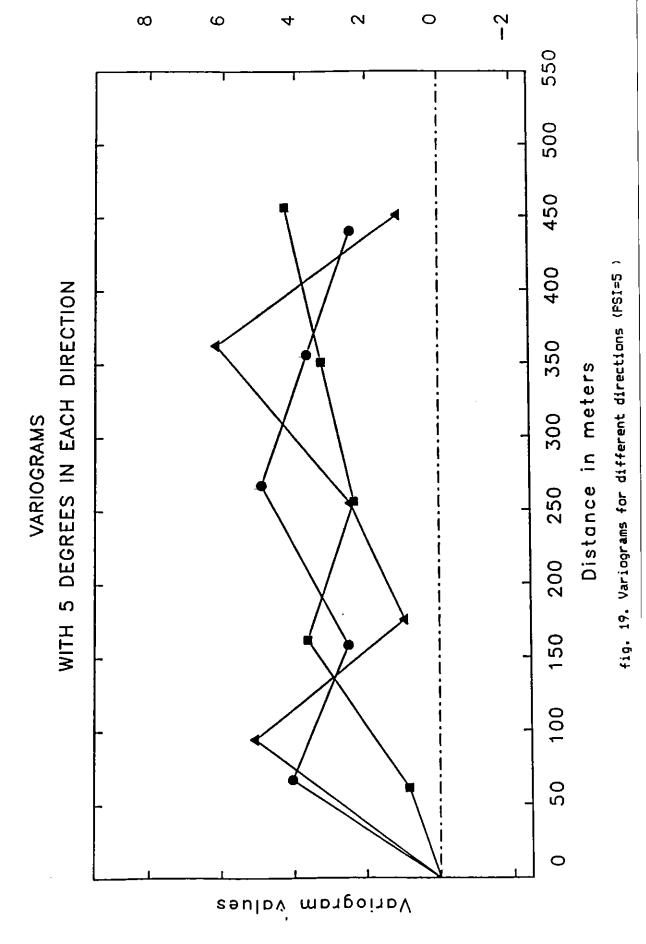
fig. 18. Variogram of the logK data set

table 11. Directional analysis (PSI=5)

VARIOGRAH . . .

STUDY OF SPATIAL VARIABILITY OF FIELD PARAMETER: HYDRAULIC CONDUCTIVITY IN H/DAY (WITH A FIELD OF 5. DEGREES IN EACH DIRECTION)

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•	400 500 500 600	20 16	179E+00	. 2452E+01 . 4394E+01	440. 6 542. 7	
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80.00	900 900 9001000	5		. 1003E+01 . 1755E+01	857. 0 7 33. 3	
99. 99		-				
-	-					
, de			V /	ARIOGRAM		
		STUDY OF SP	ATIAL VARIABILITY (CONDUCTIVITY IN MA	OF FIELD PARAMETER: /DAY 5 DECREES IN EA		
		HIBRHOLIO (HITH A FIELD OF	5. DEGREES IN EA	CH DIRECTION >	
	់ង្វ				-	LOCK
f i		40000.00				
STEP IN ME	· '=	, 1000E+03				
UPPER LIM		. 2689E+01				45.
OENERAL M	EAN OF Z =	. 4913E+00				
OENERAL V	ARIANCE OF Z =	. 3062E+01				
DENERAL SI	KEHNESS OF Z =	6965E+00			•	
GENERAL KI	URTOSIS OF Z =	. 2339E+01			AVERAGE DISTANCE	
51.	DISTANCE IN MET	ER NO. OF	PAIRS DRIFT	VARIOGRAM	AVERAGE DISTRICE	
	0 100 '	1	, 130E+01	. 6513E+00	61. 2 162. 1	
	100 200 200 300	1 <u>0</u>		, 3641E+01 , 2356E+01	256. 2	
_	300 400	14	. 106E+00	. 3262E+01 . 4298E+01	351, 2 456, 8	
	400 500 500 600	1 2	634E+00	. 3401E+01 . 2573E+01	559. 1 640. 0	
:	600 700 700 800	9 7	-, 883E+00	3083E+01 2231E+01	735. 8 838. 4	
	900 900 9001000	3	207E+00	, 9569E+00	931. 1 1020. B	
99. 99	10001100	1	. 307E+00	. 4698 E-01	1020.	
No.		- · ·	· ·	ARIDORAH	•	
			•	4 X 1 U V X A N	.•	
		STUDY OF SP	ATIAL VARIABILITY	DE FIELD PARAMETER:	•	
		HYDRAULIC	CONDUCTIVITY IN M. HITH A FIELD OF	5. DEGREES IN EA	CH DIRECTION)	
						LOCK
STEP IN ME	TER =	. 1000E+03				
UPPER LIMI	T FOR Z -	. 2689E+01				
CENERAL ME	AN DF Z =	. 4913E+00				. 90.
GENERAL VA	RIANCE OF Z =	. 3062E+01				
CENERAL SK	EWNESS OF Z =	6965E+00		ı		
GENERAL KU	RTOSIS OF Z =	. 2337E+01		·		
	DISTANCE IN MET	ER NO OF	PAIRS DRIFT	VARIOGRAM	AVERAGE DISTANCE	
				\$85.4F - 8 -	6. 1	
	0 100	1 9	31 9 E+01 . 577E+00	. 5084E+01 . 9612E+00	94. 6 175. B 255. 0	
	200 300	15	. 879E-01	. 2456E+01	233 O	



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VARIOCRAM

STUDY OF SPATIAL VARIABILITY OF FIELD PARAMETER: HYDRAULIC CONDUCTIVITY IN M/DAY (WITH A FIELD OF 20 DEGREES IN EACH DIRECTION)

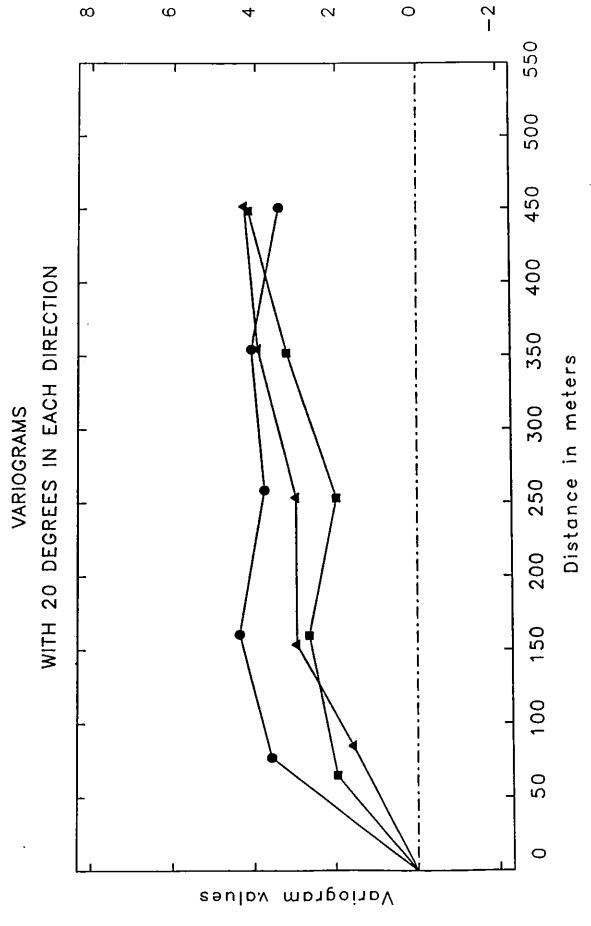
					LDCK
6TEP IN METER =	. 1000E+03				
UPPER LIMIT FOR Z =	. 2687E+01				
GENERAL MEAN OF Z =	. 4913E+00				. o.
GENERAL VARIANCE OF Z	. 3062E+01				
	6965E+00				
GENERAL KURTOSIS OF Z	, 2339E+01				
DISTANCE IN MET	ER NO. OF PAIRS	DRIFT	VARIOGRAM	AVERAGE DISTANCE	
0 100 100 200 200 300 300 400 400 500 500 600 600 700 700 800 800 900 900 1000 1000 1100	20 45 58 65 64 59 59 59 40 32 13	- 866E-01 .508E+00 .43E+00 - 154E+00 - 268E+00 - 494E+00 .868E-01 .198E+00 - 301E+00 .301E+00 .184E+01	3886E+01 4371E+01 3761E+01 4080E+01 3410E+01 3708E+01 1886E+01 1886E+01 2801E+01 2886E+01	77. 1 160. 6 238. 6 354. 6 451. 0 343. 8 653. 2 748. 7 850. 5 937. 0 1005. 3	
ŗ					
	STUDY OF SPATIAL V HYDRAULIC CONDUC (WITH	ARIABILITY OF F	I D G R A M IELD PARAMETER: DEOREES IN EACH		
STEP IN METER	MYDRAULIC CONDUC	ARIABILITY OF F	IELD PARAMETER:	DIRECTION)	LOGK
STEP IN METER =	HYDRAULIC COMDUC (WITH	ARIABILITY OF F	IELD PARAMETER:		LOGK
•	HYDRAULIC CONDUC (HITH . 1000€+03	ARIABILITY OF F	IELD PARAMETER:		·
UPPER LIHIT FOR Z =	HYDRAULIC CONDUC (WITH . 1000E+03 . 2689E+01	ARIABILITY OF F	IELD PARAMETER:		:
UPPER LIMIT FOR Z = OENERAL MEAN OF Z =	HYDRAULIC CONDUC (HITH . 1000E+03 . 2689E+01 . 4913E+00	ARIABILITY OF F	IELD PARAMETER:		
UPPER LIMIT FOR Z = QENERAL MEAN OF Z = QENERAL VARIANCE OF Z =	HYDRAULIC CONDUC (WITH . 1000E+03 . 2689E+01 . 4913E+00 . 3062E+01	ARIABILITY OF F	IELD PARAMETER:		
UPPER LIMIT FOR Z = GENERAL MEAN OF Z = GENERAL VARIANCE DF Z = GENERAL SKEWNESS OF Z =	HYDRAULIC CONDOX (WITH . 1000E+03 . 2689E+01 . 4913E+00 . 3062E+01 ~. 6965E+00 . 2339E+01	ARIABILITY OF F	IELD PARAMETER:		45.

VARIDERAM

STUDY OF SRATIAL VARIABILITY OF FIELD PARAMETER: HYDRAULIC CONDUCTIVITY IN M/DAY (WITH A FIELD OF 20. DEGREES IN EACH DIRECTION)

			LOCK
STEP IN METER = . 1000E+03			· · · · · · · · · · · · · · · · · · ·
UPPER LIMIT FOR Z ≈ .2689E+01			
GENERAL MEAN OF Z = .4913E+00			. 90 .
GENERAL VARIANCE OF Z = .3062E+01			
GENERAL SKEWNESS OF Z = 6965E+00			
GENERAL KURTOSIS OF Z = .2339E+01			
DISTANCE IN METER NO. OF PAIRS	DRIFT	VARIDGRAM	AVERAGE DISTANCE
0 100 11 100 200 24 200 300 48 300 400 51 400 500 41 500 600 48 600 700 37 700 800 10	- 325E+00 - 127E+00 - 118E+00 - 161E-01 - 268E+00 - 334E-01 - 834E+00 - 855E+00	1572E+01 2970E+01 2985E+01 3928E+01 4270E+01 3299E+01 4114E+01 1519E+01	85.2 153.7 253.6 354.8 451.9 548.5 648.6 730.3
99. 9 9			

99. 99



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fig. 20. Variograms for different directions (PSI=20)

every direction. This is due to the fact that the number of observations is extremely low. In the second graph, which represents more observations, there are less irregularities, and there does not seem to appear any influence of the direction. Thus we may say that the parameter (logK) shows isotropic properties. This was found to be true in the previous studies too.

4.7.2. Search for observational errors.

Through the "maxvar" option of the program VARIO1, a search for observational errors was tried. The idea was that, if a point appeared at the same time at the "maxvar pair" of many classes, it could carry an observational error in it.

In the variogram of logK values (table 10), we observe that the points 6,55,73 appear three times each in the maxvar pairs. The above-mentioned points, happen to have very low or very high values of the parameter, which could alone explain the fact that they appeared at the maxvar pairs. Nevertheless, we construct the variogram of the remaining observations (table 13). Its plott is reffered to as "second variogram" in figure 21, where it can be seen that the variogram, having not changed in shape, just shifted towards the distance axis, showing lower sill or sample variance.

We then calculate a "third" variogram (table 14), after having excluded the point number 25, which appears 4 times in the

table 13. Variogram excluding points 6,55,73

VARIOGRAM

STUDY OF SPATIAL VARIABILITY OF FIELD PARAMETER:
HYDRAULIC CONDUCTIVITY IN M/DAY
HYDRAULIC CONDUCTIVITY IN M/DAY
HYDRAULIC CONDUCTIVITY OF 180. DEGREES IN EACH DIRECTION

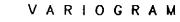
	Xeon		2				MATUAR PAIR		224-0888842 	
EACH DIRECTION /		• •	•	•	•		TOTAL DIGITARY	AVERME DISIENCE	76. 133. 231. 24. 24. 24. 26. 26. 26. 26. 26. 26. 26. 26. 26. 26	
DEGREES IN								VARIDORAH	24466 30716601 22146601 22746601 22776601 1876601 13096601	
A FIELD OF 180.								DRIFT	11.1.1.1.1.1.1.1.1.1.1.1.1.1.1.1.1.1.1	
HIIM		. 1000E+03	. 2689E+01	. 3678E+00	. 2703E+01	6673E+00	. 2304E+01	R NO. OF PAIRS	1192448991 1192448991 129391	
		•		•	OF 2 =	z	OF 2 =	DISTANCE IN METER		•
		STEP IN METER	UPPER LIMIT FOR Z	CENERAL MEAN OF Z	GENERAL VARIANCE OF	GENERAL SKEWNESS OF	GENERAL KURTOSIS OF 2	DISTA	60000000000000000000000000000000000000	

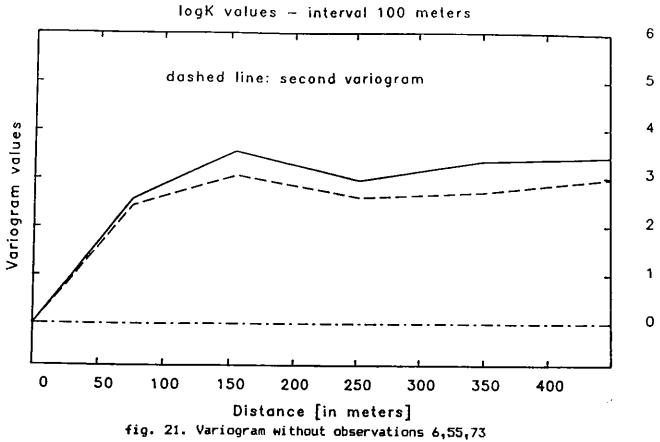
table 14. Variogram excluding points 6,26,55,73

VARIOGRAM

STUDY OF SPATIAL VARIABILITY OF FIELD PARAMETER: HYDRAULIC CONDUCTIVITY IN M/DAY (WITH A FIELD OF 180. DEGREES IN EACH DIRECTION

	TOOK			i			MAXVAR PAIR	224448588528
•	•••	:	:.	•	:		AVERAGE DISTANCE	1,76 2,51 2,51 2,51 2,51 2,51 1,51 1,51 1,51
							VARIDORAM	2391E+01 2383E+01 2382E+01 2832E+01 2804E+01 2914E+01 2317E+01 1309E+01
							DRIFT	1.1.1.1.1.1.1.1.1.1.1.1.1.1.1.1.1.1.1.
	1000€+03	. 2689E+01	. 6130E+00	. 2380E+01	6798E+00	2359E+01	NO. OF PAIRS	1289 945 945 945 945 945 945 945 945 945 94
	STEP IN METER 16	UPPER LIMIT FOR 2 . 2	GENERAL MEAN OF Z	GENERAL VARIANCE DF Z =	CENERAL SKEWNESS OF Z =	GENERAL KURTOSIS OF Z =	DISTANCE IN METER	% .% 2000 2000 2000 2000 2000 2000 2000 20
	STE	345		SER	S.	CER		\$





VARIOGRAM

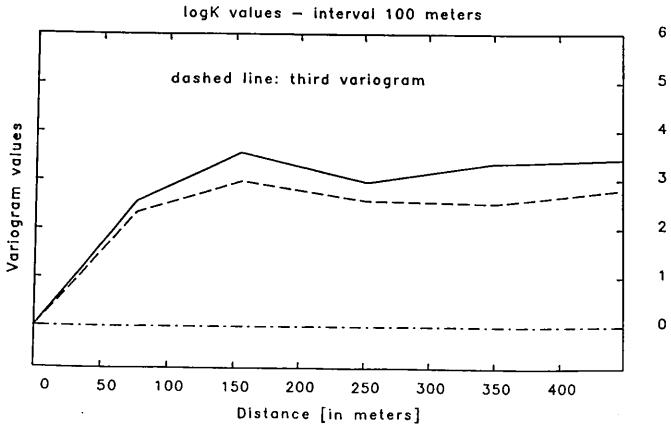


fig. 22. Variogram without observations 6,26,55.73

maxvar pairs of the second one. The resulting plott (figure 22), is very similar to the second one, showing that no observational error was detected.

4.7.3. Conclusion.

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Considering the two previous conclusions, that is, isotropism and no observational error detection, we interprete the variograms in figures 17 and 18. Both of them, show a pure nugget effect, which can be explained by the fact that, variations in the data exist at a scale smaller than the sampling distances.

4.7.4. Comparison of the variogram with previous studies.

The variogram of the present study (logK values), was plotted together with the variograms of the two previous studies (figure 23). There appear to be a lot of differences in the structure of the three variograms. The effect of the plot size, is responsible for these differences.

4.7.5. Total variogram.

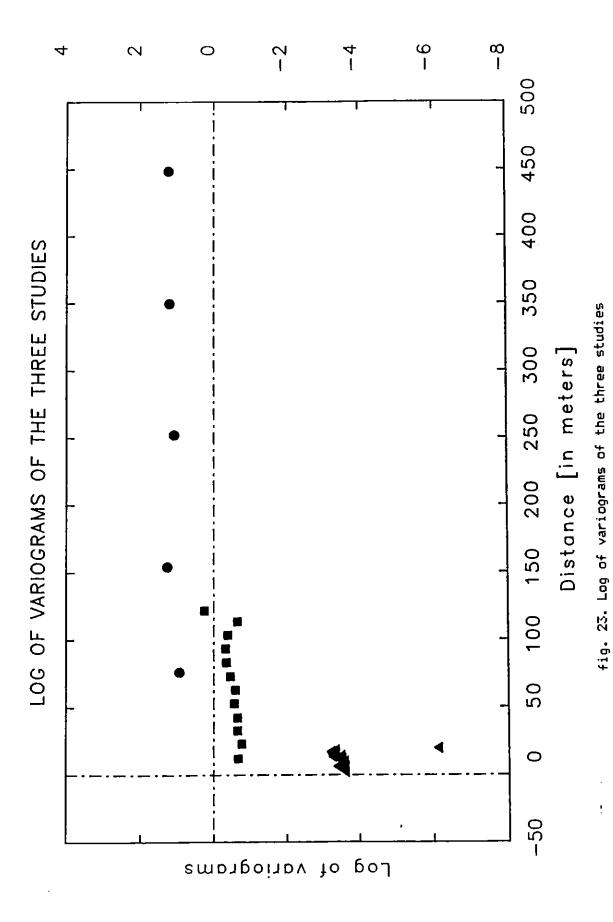
The coordinates of the points of observation of the two previous studies, were transformed into the system of the present one. Two points of the study of Tan were dropped, since they corresponded to 0 conductivity values. The rest, a total of 242 points, are shown in figure 24.

The total variogram is then calculated, for the 242

points (table 15). The results are plotted in fig. 25. Apart from the points, the line of the proposed model for the variogram is plotted there. It is a model with a sill of $C_1 = 3.79$ and a range of a=142.9m. Since there are no points between, the model is increasing linearly until the point (142.9, 3.79). This model will be used in the application of the kriging estimation technique in chapter 6.

4.7.6. One-point variograms of each study.

Considering the points of each study separately, as belonging to the same class, we calculate one variogram value for each study. The results are situated in table 16, and plotted in fig. 26. No model variogram can be fitted to the three points, since the variogram value of the first study seemes relatively very small with respect to the average distance.



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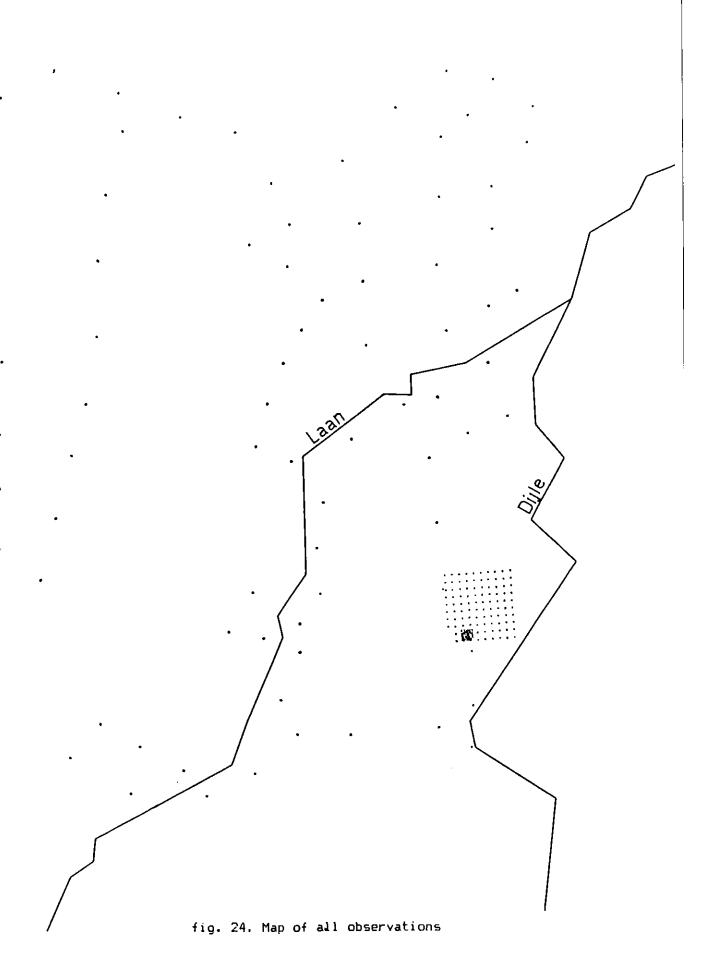


table 15. The total variogram

VARIOORAM

STUDY OF SPATIAL VARIABILITY OF FIELD PARAMETER HYDRAULIC CONDUCTIVITY IN M/DAY (WITH A FIELD OF 180. DEGREEB IN EACH DIRECTION)

LDGK			45.				MAXVAR PAIR	287747588
•••	•	••		•			AVERAGE DISTANCE	44 1994 94 46 44 44 44 44 44 44 44 44 44 44 44 44
							VARIDORAM	1173E+01 3772E+01 3772E+01 378F+01 3714E+01 376E+01 4407E+01 177E+01
							DRIFT	7038 1.6738 1.6778 1.4408 1.1988 1.1988 1.1088 1.1088 1.1088 1.1088
	. 1000E+03	. 2689E+01	7318E+00	. 2514E+01	. 2998E+00	. 2429E+01	NO. OF PAIRS	133 964 2003 2003 2003 173 173 862 321 33
	STEP IN METER . 1	UPPER LIMIT FOR Z2	CENERAL MEAN OF Z	OENERAL VARIANCE OF Z	CENERAL SKEWNESS OF Z	GENERAL KURTOSIS OF Z	DISTANCE IN METER	100 200 200 200 300 400 400 500 500 500 500 500 1000 1000
	STEP	UPPER	CENER	OENER	CENER	CENER		& · &

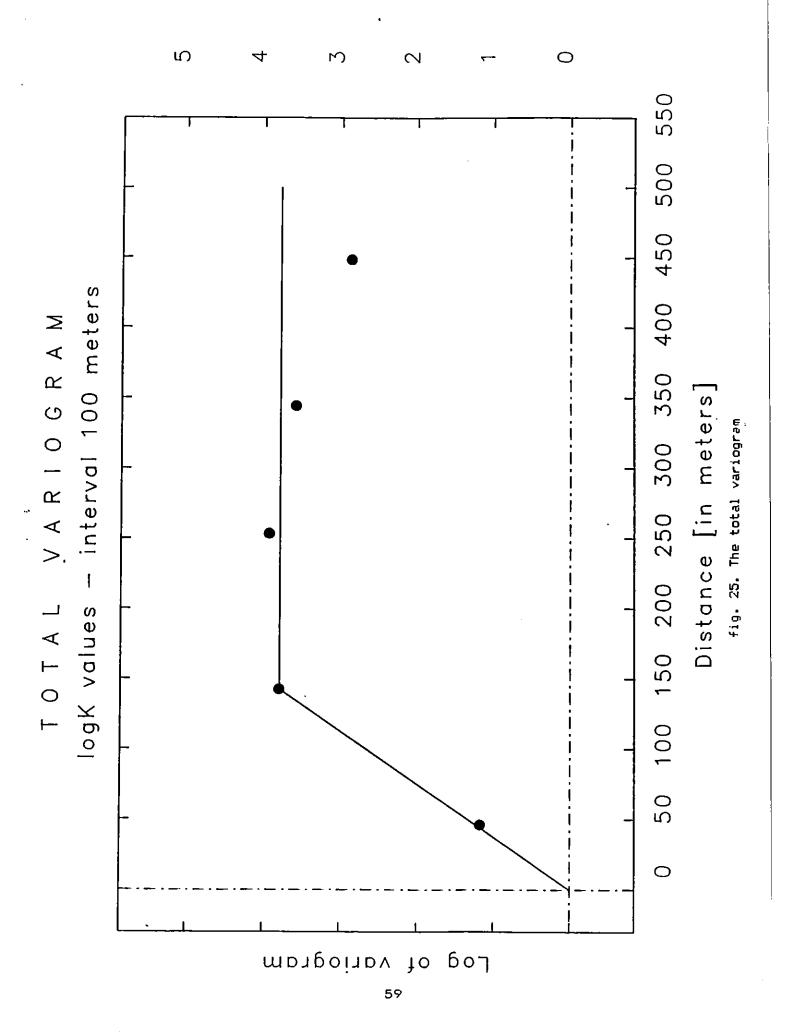


table 16. One point variograms for each study

V A R 1 0 C F A M

STUDY OF SPATIAL VAPIADILITY (HT FIELD PARAMETER CHYDRAULIC (DN)UCTIVITY K-M/DAY) (WITH A FIELD OF 180 DEGREES IN EACH DIRECTION)

			•			FDC+
STEP IN HEIER	₹ .20	000f.+02				
UPPER LIMIT FOR 2	= .10	1.390 +60				
GENERAL HEAN OF Z	٠. ۽	2451E+00				45
CENTRAL VAFIANCE OF	2	280Pt - 01				
MENERAL SKEWRESS OF	7	4493E+00				
CENERAL KURTOSIS DE S	7 = 3	2327£ +01				
DISTANCE	IN METER	NO OF PAIRS	DRIFT	VARIOGRAM	AVERAGE DISTANCE	
99.99	- 20	1953	. 259E-01	. 27226-01	8. 4	'

VARIDORAH

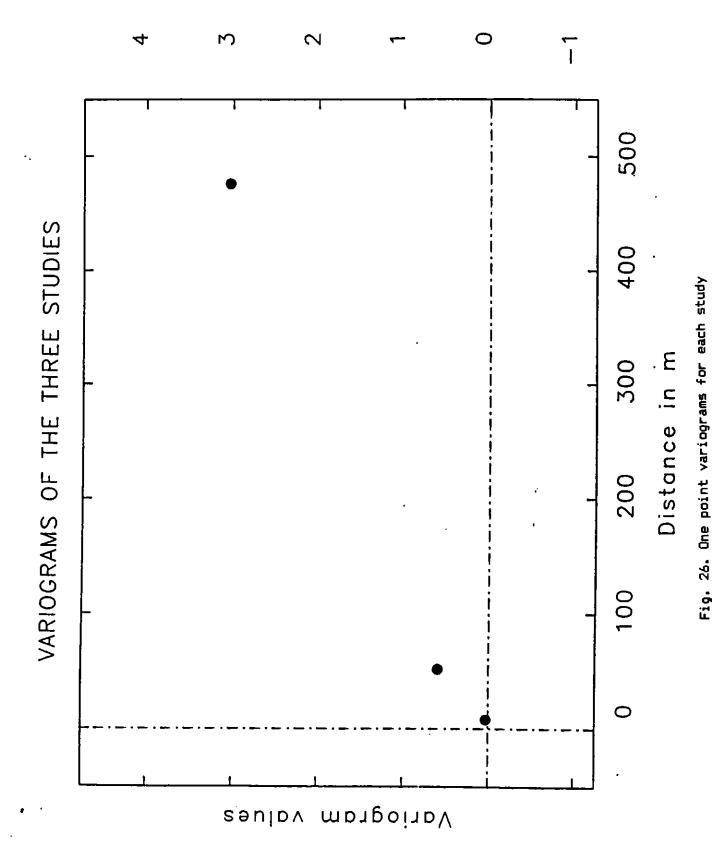
STUDY OF SPATIAL WARIABILITY OF FIELD PARAMETER (HYDRAULIC CONDUCTIVITY K-M/DAY) (WITH A FIELD OF 180. DEGREES IN EACH DIRECTION)

, ,						LOOK
∯STEP #	N METER	. 2000E+03	•			•
UPPER I	INIT FOR Z	3413E+00				
POENERAL	. HEAN OF Z	2007E+01				45.
DENERAL	VARIANCE OF Z	6636E+00	•			
DENERAL	. SKEHNESS OF Z	9873E-01			•	
OENERAL	. KURTOSIS OF Z	3975E+01				
44	DISTANCE IN	METER NO. OF PAIRS	DRIFT	VARIOGRAM	AVERAGE DISTANCE	
79. 9 9	, o —— æ	00· 4753	-, 128E+00	. 5926E+00	51.7	

VARIDORAM

STUDY OF SPATIAL VARIABILITY OF FIELD PARAMETER: HYDRAULIC CONDUCTIVITY IN H/DAY (WITH A FIELD OF 180. DEGREES IN EACH DIRECTION)

02000 3081	- 200E+00	, 3039E+01	475. B
DISTANCE IN METER NO. OF PAI	RS DRIFT	VARIDORAH	AVERAGE DISTANCE
OENERAL KURTOSIS OF Z = .2339E+01			
GENERAL EKENNESS OF Z =6965E+00			
DENERAL VARIANCE OF Z			
GENERAL MEAN OF Z = .4913E+00			
UPPER LIMIT FOR Z = .2687E+01			. 45.
STEP IN METER = ,2000E+04			
			; LD0+
•			



4.8. Kolmogorov-Smirnov goodness of fit test.

Looking for the distribution of the data sets, we will test the hypothesis that it is normal. In order to test this hypothesis, we will use the Kolmogorov-Smirnov goodness of fit test, which, it should be stressed, cannot be applied in cases where the data is correllated.

The theoretical probabilities of an observation being smaller or equal to each specific one (under the hypothesis of normal distribution), are first calculated. This is done by first normalising the values with the formula:

$$Z=(K-m)/s$$
 (4.20)

where m is the mean, and s the standard deviation, and second, reading the probability from the standard normal curve table:

$$F(K)=P(Z(z)) \qquad (4.21)$$

The corresponding experimental probabilities are then also calculated. The observed values have to be in ascending order. Then, as the experimental probability of a measurement being smaller or equal to the Ith observation with value K, we consider:

$$Fe(K)=(I-0.5)/n$$
 (4.22)

where n is the total number of observations.

For each specific observation K, the difference between the experimental probability is calculated:

$$di=Fe(Ki)-F(Ki)$$
 (4.23)

The greatest value of the dis is then compared with standard

functions, depending upon the significance level considered: for significance level 5%, dmax allowed is $(1.36/\sqrt{n})$ for significance level 10%, dmax allowed is $(1.22/\sqrt{n})$

4.9. Application of the Kolmogorov-Smirnov test.

the present study, the distances between the observation points are of the order of magnitude of 50 to In order to apply the goodness of fit test for normality of the observations, we have first to ensure the correllatedness among them. The model of the total variogram, has a range of 142m. So, to ensure the non-correlatedness we would have to exclude all points with distances smaller than 140m. in that case, we would have been left with very few observations. For this reason, we select 35 observations between which there is distance no smaller than 90 meters. since hydraulic conductivities of points with smaller distances, be can considered as being correllated.

A map of these points is situated in the following page. The program VARIO3 is then used to make the test, and the results can be seen in table 18 for the logK values, and in table 19 for the K values. Figure 28 shows the theoretical and experimental probabilities versus the logK values, and figure 29 versus the K values. The results show that both the normality of logK and K values are accepted at a significance level of 5%, and rejected at a level of 10%.

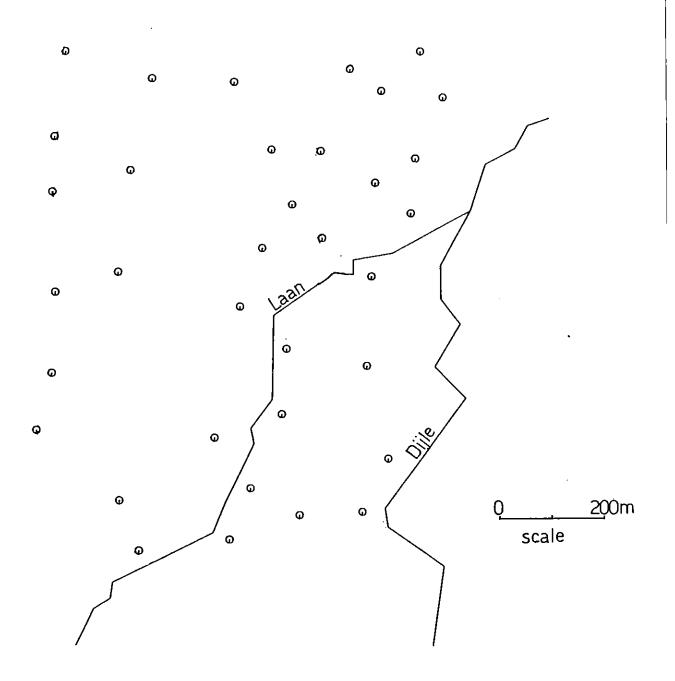


fig. 27. Observations considered for the K-S test

table 17. Kolmogorov-Smirnov test for logK values

NUMBER OF SAMPLES = 35

RA123456789012345678901234 11123456789012345678901234	-3. 456 -3. 071 -1. 328 -1. 328 -1. 328 -1. 228 -1. 229 -1. 428 -1. 428 -1. 428 -1. 428 -1. 429 1. 1205 1. 205 1. 305 1. 305 1. 7436 1. 743	NDR. 427 -1. 327 -1. 327 -1. 327 -1. 3257 -1. 3257 -1. 9776 -1. 9776 -1. 3257 -1. 9776 -1. 3257 -1. 32	EXPT. PROB . 0142857 . 0428571 . 0714286 . 1000000 . 1285714 . 1571429 . 1857143 . 2142857 . 2428571 . 2714286 . 3000000 . 3285714 . 3571429 . 3857143 . 4142857 . 4428571 . 4714286 . 5000000 . 5285714 . 5571429 . 5857143 . 6142857 . 6428571 . 6714286 . 7000000 . 7285714 . 7571429 . 7857143 . 8142857 . 8428571 . 8714286 . 7000000 . 7285714 . 7571429 . 7857143 . 8142857 . 8428571 . 8714286 . 9000000	THEO. PROB . 0043696 . 0085033 . 0789338 . 0789338 . 0931946 . 0956069 . 1041386 . 1420823 . 1772225 . 2202541 . 2578152 . 3315578 . 3669936 . 5149216 . 5765070 . 5889165 . 6133477 . 6317391 . 6373160 . 6659252 . 7042528 . 7107422 . 7565202 . 7569965 . 7649221 . 7730932 . 7906138 . 8047884 . 8144588 . 8193256 . 8383119	D1 010 034 008 007 033 047 038 014 038 158 175 170 160 137 145 196 015 016 017 017 017 018 019 019 019 019 019 019 019 019	D2 . 000 . 036 . 022 . 004 . 024 . 015 . 008 . 047 . 186 . 219 . 166 . 174 . 166 . 174 . 147 . 125 . 131 . 114 . 086 . 045 . 0	D . 010 . 034 . 032 . 033 . 047 . 023 . 047 . 023 . 047 . 047 . 189 . 166 . 174 . 147 . 125 . 114 . 086 . 045 . 031 . 037 . 081 . 057 . 081 . 059
34 35	2. 450 2. 563 2. 689	. 988 1. 057 1. 133	. 9285714 . 9571429 . 9857143	. 8383119 . 8546852 . 8714897	. 0 7 0 . 102 . 114	. 062 . 074 . 086	. 0 90 . 102 . 114

DMAX = .219

THE NORMALITY OF THE OBSERVATIONS IS ACCEPTED FOR SIGNIFICANCE LEVEL 5% SINCE DMAX= .219 AND THE LIMIT IS .230

THE NORMALITY OF THE OBSERVATIONS IS REJECTED FOR SIGNIFICANCE LEVEL 10% SINCE DMAX= .219 AND THE LIMIT IS .206

THE MEAN IS . 834177 AND STD. IS 1.636142

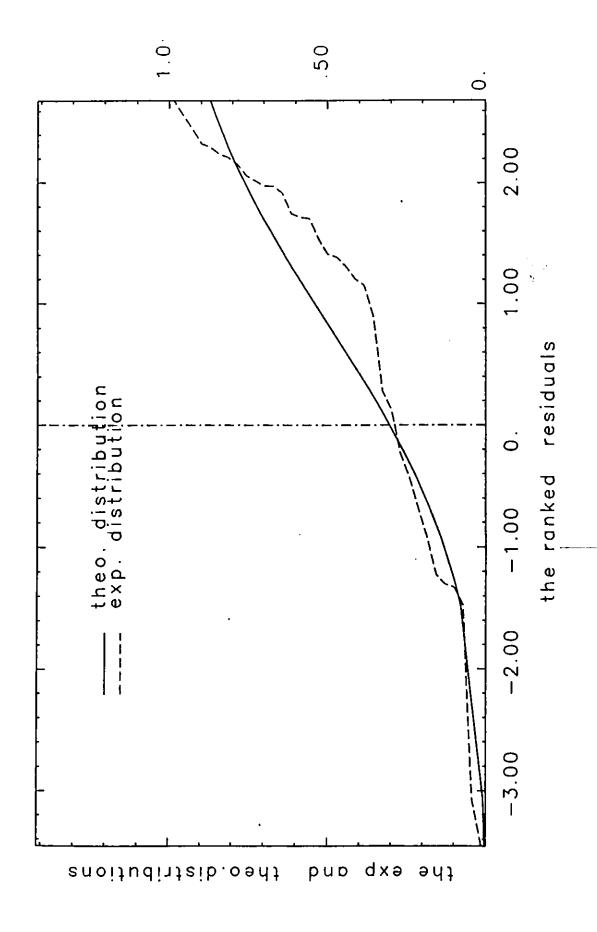


Fig. 28. Distribution curves for log of K values.

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table 18. Kolmogorov-Smirnov test for K values

NUMBER OF SAMPLES= 35

RA 12345678901234567890123456	Z VAL . 0001 . 0033 . 050 . 1207 . 1207 . 1373 1. 859 14. 919 15. 204 24. 6326 551. 353 851. 3	NOR. VAL 703 703 703 703 7003 7001 7001 692 685 564 580 564 402 402 402 2016 114 2016 119	EXPT. PROB . 0142857 . 0428571 . 0714286 .1000000 .1285714 .1571429 .1857143 .2142857 .2428571 .2714286 .3000000 .3285714 .3571429 .3857143 .4142857 .442857 .442857 .442857 .442857 .442857 .4714286 .5000000 .5285714 .5571429 .5857143 .6142857 .6428571 .6714286 .7000000	THEO. PROB . 2407369 . 2407382 . 2410267 . 2410638 . 2410709 . 2410781 . 2412644 . 2415033 . 2417525 . 2425421 . 2445475 . 2461260 . 2628162 . 2807511 . 2962641 . 2971280 . 3114929 . 3158058 . 3432227 . 3766052 . 4000286 . 4128792 . 5062157 . 5454747 . 5473564 . 5800710	D1 227 170 1412 084 0527 029 0055 0894 1058 1440 185 1861 1861 1861 1861 1861 187 187	D2 .000 .227 .178 .170 .141 .113 .084 .028 .027 .027 .034 .066 .074 .074 .131 .131 .137 .137 .137 .137 .137	D72278 . 1770 . 1413 . 0846 . 0029 . 0055 . 084 . 185 . 1661 . 185 . 147 . 153 . 149
22345 2245 2254 2267 227 237 237 237 237 237 237 237 237 23	55. 353 82. 351 93. 458	220 . 016 . 114	. 6142857 . 6428571 . 6714286	. 4128792 . 5062157 . 5454947	. 201 . 137 . 126 . 153	. 173 . 108 . 097 . 124	. 201 . 137 . 126 . 153

DMAX = . 227

THE NORMALITY OF THE OBSERVATIONS IS ACCEPTED FOR SIGNIFICANCE LEVEL 5% SINCE DMAX= .227 AND THE LIMIT IS .230

THE NORMALITY OF THE OBSERVATIONS IS REJECTED FOR SIGNIFICANCE LEVEL 10% SINCE DMAX= .227 AND THE LIMIT IS .206

THE MEAN IS 80. 566149 AND STD. IS 114. 555251

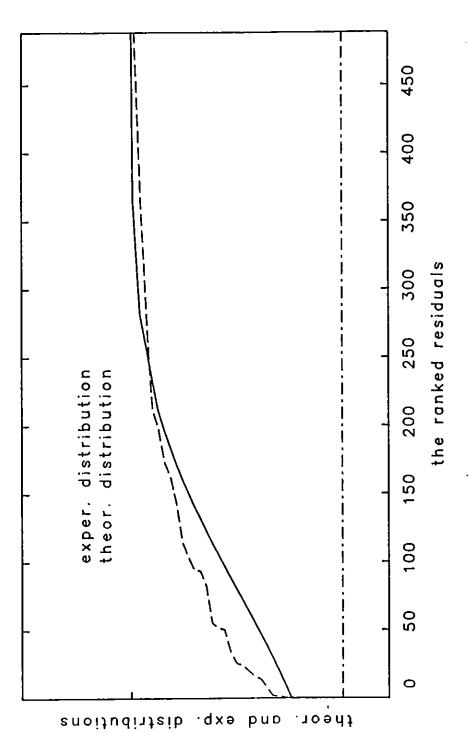


Fig. 29. Distribution curves for K values